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Supplemental Material

Figure S1. Inverse-distance-weighted (IDW) root mean square error (RMSE) values for different combinations of n and p, where n is the number of neighboring temperature values included in the weighted average estimate, and p is the weight or power parameter applied to each neighbor. The combination of n and p that resulted in the lowest RMSE were chosen for the final IDW calculations reported in this study. Note that a jitter factor was applied to the x-axis values of data to improve readability.

Figure S2. Latitudinal MAT transects and associated uncertainties from the Monte Carlo simulations, for the Campanian (A), early Maastrichtian (B), and late Maastrichtian (C) analysis windows. The solid lines show the mean MAT values for the western edge (red), center (blue), and eastern edge (green) of the study areas. The shaded regions show the maximum and minimum range of the 1000 unique interpolated temperature maps.

Figure S3. Raw fossil pollen (A, B, C) and leaf (D, E, F) bar charts showing differences in north/south ecoregion abundances.

Figure S4. Spatial patterns in fossil pollen (A-C) and leaf (D-F) ecoregion abundances. Abundance categories I, II, III, and IV indicate the mean, minimum, and minimum interpolated abundances for each ecoregion type.

DETAILED METHODS AND MATERIALS 1. Temperature estimate uncertainties

The uncertainties associated with each individual MAT estimate were combined into a total MAT uncertainty using a root sum of squares method:

 $u_c(T) = \sqrt{\Sigma_i(u(T_i))^2}$ (1) where $u_c(T)$ is the combined standard uncertainty, and $u(T_i)^2$ is the uncertainty associated with the individual MAT estimate T_i .

2. Raw versus adjusted temperature estimates

Three analysis windows were chosen with the goal of minimizing the amount of time averaging imposed on our interpolations, while still retaining sufficient (at least 10) unique sample localities for the spatial interpolation calculations (Table S1). Because two of these analysis windows (Campanian and late Maastrichtian) still covered more than 2 million years, we used the global Cretaceous temperature curves of Friedrich et al. (2012) and O'Brien et al. (2017) to adjust our temperature data to account for secular changes to global MAT. The mean difference between the interpolations using the raw temperature data versus the adjusted temperature data is -1 °C (-1.5 to -0.4 °C) for the Campanian analysis window, -0.1 °C (-0.67-0.56 °C) for the early Maastrichtian analysis window, and -0.76 °C (-1.5 to -0.4) for the late Maastrichtian analysis window. Because the difference between the raw and adjusted temperatures are small and within the range of uncertainty, we report only the raw temperature interpolations in this study.

3. Spatial interpolation methods

Spatially interpolated maps of MAT were created for each analysis window using the inverse distance weighted (IDW) method (Shepard, 1968) as implemented in the custom MATLAB function gIDW (Langella, 2020) which estimates a value (e.g., MAT) for an unsampled location based on the weighted average of surrounding, sampled points using the following equation:

$$T_j = \frac{\sum_{i=1}^{n} \frac{T_i}{d_{ij}^p}}{\sum_{i=1}^{n} \left(\frac{1}{d_{ij}^p}\right)} (2)$$

where T_j is the estimated MAT at unsampled location j, T_i is a known MAT at sampled location i, d_{ij} is the distance between locations j and i, n is the number of sampled locations or neighbors included in the weighted average, and p is the weight or power parameter applied to d_{ij} as an exponent such that as d_{ij} increases in magnitude, the effect of T_i on T_j decreases. Using this equation, the final estimated value of T_j can be adjusted by changing p and n.

Root mean square errors (RMSE) were calculated for the interpolations by using a jackknife resampling approach. This involved removing one sample from our paleotemperature data set, performing the IDW interpolation with the remaining samples, and then comparing the estimated temperature at the location of the removed sample to the actual reconstructed value. This process was then repeated for each sample in the paleotemperature data set. Final p and n values were then chosen to minimize RMSE for each analysis window (Fig. S1).

We chose to use the relatively simple IDW method rather than a more complex statistical interpolation method such as Kriging for two reasons. First, the structural analysis of our temperature data sets did not produce semivariogram results that fit well with any commonly used semivariogram models (e.g., circular, spherical, exponential, linear, etc.). Second, the Kriging results were consistent with the IDW results, and did not provide any additional information.

4. Identifying MAT transition zones

For the purposes of this study, we define a transition zone as a contiguous latitudinal zone with a mean slope that is significantly more negative than the mean Late Cretaceous latitudinal temperature gradient (-0.3 to -0.4 °C °lat⁻¹, 4). For each of the 1000 Monte Carlo IDW iterations calculated for the three MAT data sets, latitudinal MAT transects were analyzed at 1° longitude intervals. The script identifies contiguous areas of negative MAT slopes along each transect where at least one cell has a slope more negative than some threshold temperature slope (see below). Cells within the identified latitudinal zone were then assigned a value of 1, while cells outside the zone were assigned a value of 0. By summing the results of this analysis for the 1000 Monte Carlo iterations, we were able to identify how often a given cell was identified as being inside the steep latitudinal gradient zone (e.g., Fig. S3). Cells that were flagged as being within the steep latitudinal gradient zone. These calculations were performed using three different slope threshold values: -0.3 °C °lat⁻¹ (minimum estimated Late Cretaceous temperature gradient), -0.4 °C °lat⁻¹ (maximum estimated Late Cretaceous temperature gradient), and -0.5 °C °lat⁻¹ (conservative temperature gradient estimate).

Latitudinal temperature gradients were calculated for each longitudinal transect in the 1000 iterations of the Campanian, early Maastrichtian, and late Maastrichtian spatial interpolations using the following multivariate linear regression model:

 $MAT = b_0 + (b_1 \times L) + (b_2 \times TZ) + (b_3 \times L:TZ) + \varepsilon (3)$

where *L* is the paleolatitude, *TZ* is an identifier variable that indicates whether a given cell is inside or outside the previously described temperature transition zone, *L*:*TZ* is an interaction variable between *L* and *TZ*, and ε is the error term associated with the linear regressions. The regression coefficients b₀, b₁, b₂, and b₃ are associated with L, TZ, and L:TZ, respectively. The interaction effect allows for two separate lines to be fit to the data, depending on whether a given cell is inside or outside the transition zone. For cells outside the transition zone, b₀ is the yintercept value and b₁ is the change in MAT for a one unit increase in paleolatitude. For cells inside the transition zone, b₀ + b₂ is the y-intercept value and b₁ + b₃ is the change in MAT for a one unit increase in paleolatitude.

5. Fossil pollen and leaf ecoregion assignment

Ecoregions simplified for paleo-applications were defined based on available information for sample localities (c.f., USDA Levels I and II). For sites where authors provided a detailed environment and/or ecosystem description, this was conserved in our analysis. For sites where no such description was provided, both depositional environment (either from the original study, or compiled from the Paleobiology Database; PBDB data downloaded from the Paleobiology Database on April 9, 2020) and floral assemblage (from the original study) were used to define an ecoregion. For floral assemblages (both pollen and leaf), an ecoregion was identified by attributing fossil taxa to a modern nearest living relative (unattributed or disputed groups assigned conservatively at higher taxonomic levels), and then linking modern distributions (Omernik and Griffith, 2014) to existing low-resolution ecoregion maps. For consistency, this method was verified against ecoregion designations for which authors provided an original description, and any discrepancies or unattributable groups are noted in Data sets S3 and S4.

6. Fossil pollen and leaf Dice Similarity analysis

Pairwise comparisons of biodiversity using the Dice Similarity Index were performed in the R package *fossil* version 0.4.0 (Vavrek and Larsson, 2010). The resulting data set contained the similarity value between each formation for which pollen and leaf data was known (each data set analyzed separately), as well as the geodesic line distance between them (also from the R package *fossil* version 0.4.0), mean age of the formation, and the time gap between each pair of formations.

7. Floral climatic boundary test

Using the shift in temperature transition zones as a guide, we tested each integer of latitude from 40 to 60 degrees as being a potential boundary zone within a two-zone biogeographical distribution of plants during the latest Cretaceous of North America. If after assessing results from each of these latitudes it is found that the pattern expected of a bizonal or transitional biotic zonation is not present, then this would be evidence that the climatic boundary observed from climate proxies did not affect the distribution of plant species.

In order to test the hypothesis that a climatic barrier caused two distinct biomes of vegetation, we filtered the similarity results data set into three subsets, 72 - 80 Ma, 69 - 78 Ma, and 66 - 75 Ma. The broad time bins for this analysis were required in order to achieve an adequate sample size for calculations, although we did limit the inclusion of pairwise similarity results to those formations that are separated by three million years or less. An advantage to using broad time bins is that results will be homogenized over a larger amount of time data, meaning that any data excursions within smaller time frames will be suppressed by data from other time frames, thereby producing smoother, more conserved, variation within the pairwise similarity indices. For each of the 20 test latitudes we categorized formations whose paleolatitude fell north or south of the hypothesized boundary into respective North or South bins. Each pairwise similarity calculation was, therefore, labeled as a North-North, North-South, or South-South comparison.

This test was designed to determine the amount of statistical variation of pairwise similarity within and between hypothesized biomes. Consider a geographic area that is divided into two latitudinally-arrayed biogeographical biomes with a small zone of transition between the biomes. Additionally, within each biome there are several localities that preserve a particular set of organisms on which similarity calculations are to be made. If one were to choose a latitude and then calculate pairwise similarity between each of the localities within the northern and southern biomes, a certain amount of statistical variation would exist among the similarity values that is contingent upon the amount of mixing occurring within the true northern and southern biomes that are incorrectly lumped together within the hypothetical northern or hypothetical southern zones currently being tested. In other words, the greater the amount of inappropriately categorized fauna/flora within a single hypothetical biome, the larger the statistical variance present within the similarity values of a particular zone. Likewise, for those pairwise comparisons that cross a biome boundary, we would expect to see greater variance in the distribution of similarity values closer to the biome boundary, because localities that are in actuality closely related but incorrectly classified in the opposite biome, will be more similar to one another and therefore decrease the overall statistical variance. In short, we ran consecutive tests of within- and between-biome variance, suggesting that the true biotic boundary—if present—would be found at the position that maximizes the difference between in-group variance (i.e., the absolute value of the difference between the northern and southern similarity variance) and between-group variance. Additionally, we posit that the boundaries of a transitional zone can be estimated by finding the latitudes where the solution to the equation (between-group variance – in-group variance) is minimized (note that this is not the same as the absolute value of the difference).

One problem with this sliding window approach is that near the extremes of the test windows we will have far more formations present in one biome versus another when calculating variance. To accommodate this inequity, we randomly sampled 100 populations from the larger of the two biomes that were equal in size to the population in the smaller biome. All values from the subsampling and the raw cross-boundary variances were estimated and plotted with LOESS. These values were then standardized to 1 for plotting.

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